# Geographic Barriers to Education in Disadvantaged Communities: Evidence from High-School Openings in Israeli Arab Localities

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Geographic Barriers to Education in Disadvantaged Communities

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- 36% of youth aged 15-17 worldwide do not attend school (UNESCO, 2018)
- In Israel, high-school dropout is more common today among Arab students, and especially among the Bedouin in the Negev (17% of ages 16-17 do not attend high-school)
- Understanding the root causes of high school dropout is of great interest to policymakers.
- One of the potential challenges to high school attendance is a geographic barrier:
  - lack of access to local school in disadvantaged communities can affect their educational success.

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- One of the potential challenges to high school attendance is a geographic barrier:
  - lack of access to local school in disadvantaged communities can affect their educational success.

- Main question: What are the short- and long-term effects of opening a high school in a disadvantaged locality?
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#### Previous Literature and Contribution

- 2 Background: Israeli Arab Education System
- ③ Data and Research Design
- 4 Empirical Methodology

#### 5 Results



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- Elementary school access (developing countries).
  - Duflo, 2001; Kazianga et al. 2013; Burde and Linden, 2013; Akresh et al, 2018; Mazumder et al. 2021.
- College access (developed countries)
  - Card, 1993; Currie and Morretti, 2003; Cameron and Taber. 2004; Carneiro et al. 2011; Nybom, 2017; Kamhöfer et al., 2018; Mountjoy, 2021.
- Secondary school access in UK, Norway and France
  - Relatively few studies:
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- My study contributes to the existing literature on *high-school openings* in three key ways:
  - Improved causal analysis (diff-in-diff). Examination of school openings over time in treated vs comparison localities.
  - First study (to the best of my knowledge) to examine also the **long-term** effects of proximity to high-school.
    - Evidence on post-secondary education, employment, earnings, marriage and fertility.
  - Focus on disadvantaged communities with low educational enrollment rates. More similar to less developed countries.
    - High-School completion rates in the treated localities were 41% prior to school openings, while today's OECD average stands at 80%.

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### Arab High-School Enrollment (Not inc. Negev-Bedouin)



← Share Completed 10th Grade ← Share Completed 12th Grade

- During the period of my study, 1972-1995, the number of Arab high schools has changed significantly, from 32 to 106.
- The most important change was in geographic coverage. The percentage of people living in places with high schools increased significantly from about 28% in 1970 to about 92% in 1995.

# Share of localities with at least one high-school (Not including Negev Bedouins)



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- Unique data about the historical Arab high schools in Israel, by locality
  - Compiled from several archival paper publications on Israeli schools and localities
- Administrative records in the CBS secure research room on all Israeli Arabs born in 1950-2000.
  - Population register: gender, year of birth, locality of residence, religion, family size. (1995+)
  - Israeli censuses of 1972 and 1983
  - CBS Israeli Education Register: years of education for the individuals and their parents.
  - Earnings and Employment (Israeli Tax Authority): Wage earners in 1983-2019. Linked to individuals in ages 33-35.
  - Juvenile criminal records (for cohorts in the Bedouin sample).

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- I exploit variation across localities and years of high-school openings to examine the effects of school proximity on individuals' future outcomes.
- Diff-in-Diff framework: Examine changes in the outcomes of high-school age cohorts, before and after the opening of a local high school, compared to comparison localities where no high-school was opened in those years.

### Opening of Arab High-Schools (Not inc. Negev-Bedouins)



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- The identification assumption is that the exact *timing* of local high-school openings across municipalities is uncorrelated with local developments of the outcomes. ( "as good as random")
- This is a nontrivial assumption. Perhaps the establishment of a school is endogenous to prior locality-specific trends.
- I find that there are two characteristics that predict high school openings:
  - Population of the locality
  - Compulsory Schooling Laws

- I will justify the identification assumption by:
  - Applying event-study models to rule out prior-trends
  - Adding municipality-population controls to the regressions.
  - Apply differential cohort effects by the year of Compulsory Schooling implementation.
  - Additional robustness checks

# Examining trends

### **Event-Analysis Graphs: Education**



B. Share completed High School

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16 15 14



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 $y_{\textit{ilt}} = \alpha + \beta \textit{HighSchoolAccess}_{\textit{lt}} + \gamma \textit{X}_{\textit{ilt}} + \delta \textit{CohortSize}_{\textit{lt}} + \lambda_{\textit{crt}} + \eta_{\textit{l}} + \varepsilon_{\textit{ilt}}$ 

- Where for individual i in birth-year *t*, in locality *I*:
  - y<sub>ilt</sub>: Outcomes
  - *HighSchoolAccess<sub>lt</sub>*: Dummy indicator for having an high-school in the locality of residence at age 17 or younger.
  - $X_{ilt}$ : Gender, years of education of parents, number of siblings, religion.
  - CohortSize<sub>It</sub> : Population size of cohort in birth-year t, in locality I. (Polynomial)
  - $\eta_I$ : Locality fixed effect.
  - $\lambda_{crt}$ : Cohort effect **or** Cohort-by-region effect (birth-year t) **or** cohort-by-region-by-year of compulsory schooling implementation effect.
  - $\epsilon_{ilt}$  :Error term (may be correlated within the same locality over time).

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• Second Specification:

$$y_{ilt} = \alpha + \sum_{\tau=13, \tau\neq 18,}^{\tau=21} \beta_{\tau} * HS_{lt}^{\tau} + \gamma X_{ist} + \delta CohortSize_{lt} + \eta_l + \lambda_{crt} + \theta I(\tau_{tl} < 13) + \mu I(\tau_{tl} > 22) + \epsilon_{ilt}$$

•  $HS_{lt}^{\tau}$ : indicator equal to 1 if the individual was in age  $\tau$  when the local high-school was opened, and 0 otherwise.



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# **Educational Outcomes**

	Completed 10th Grade				
High-school access	$(1) \\ 0.100^{***} \\ (0.021)$	(2) $0.103^{***}$ (0.021)	$(3) \\ 0.075^{***} \\ (0.017)$	$(4) \\ 0.069^{***} \\ (0.016)$	$(5) \\ 0.070^{***} \\ (0.019)$
Mean outcome (age 18 in year of HS opening)	$\left] 0.577 ight $	0.577	$\left] 0.577 \right]$	0.577	0.577
Individual controls		X	Х	Х	Х
Locality's population (cohort size)			Х	Х	Х
Locality fixed-effect	Х	Х	X	Х	Х
Type of cohort fixed-effect	$\begin{array}{c} { m Cohort} { m FE} \end{array}$	Cohort FE	Cohort FE	Cohort×reg FE	ion Cohort× region × compulsory schooling year FE
Number of localities	89	89	89	89	89
Number of observations	$316,\!322$	316, 322	$316,\!322$	$316,\!322$	316,322

• Preferred estimates show a 7 percentage points (12%)  $\uparrow$  in the probability of completing 10th grade.

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	Completed High-School (12th Grade)				
High-school access	$(1) \\ 0.060^{***} \\ (0.015)$	$(2) \\ 0.073^{***} \\ (0.017)$	$(3) \\ 0.057^{***} \\ (0.015)$	$(4) \\ 0.054^{***} \\ (0.013)$	(5) $0.055^{***}$ (0.015)
Mean outcome (age 18 in year of HS opening)	0.408	0.408	0.408	0.408	0.408
Individual controls		Х	Х	Х	Х
Locality's population (cohort size)			Х	Х	Х
Locality fixed-effect	Х	Х	X	Х	Х
Type of cohort fixed-effect	$\begin{array}{c} { m Cohort} { m FE} \end{array}$	$\begin{array}{c} { m Cohort} { m FE} \end{array}$	Cohort FE	Cohort×reg FE	tion Cohort× region × compulsory schooling year FE
Number of localities	89	89	89	89	89
Number of observations	316,322	$316,\!322$	$316,\!322$	316,322	316,322

• Preferred estimates show a 5.5 percentage points (13%)  $\uparrow$  in the probability of completing high-school.

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# Results: Education, by gender

<b>Dependent Variable</b> Completed 10th grade	$\begin{array}{c} \textbf{All} \\ 0.070^{***} \\ (0.019) \\ \text{Mean}{=}0.577 \end{array}$	$\begin{array}{c} {\bf Boys} \\ 0.061^{***} \\ (0.018) \\ {\rm Mean}{=}0.648 \end{array}$	$\begin{array}{c} {\bf Girls} \\ 0.080^{***} \\ (0.023) \\ {\rm Mean}{=}0.504 \end{array}$
Completed 11th grade	$0.054^{***}$	$0.039^{***}$	$0.072^{***}$
	(0.015)	(0.015)	(0.019)
	Mean= $0.449$	Mean= $0.512$	Mean=0.382
Completed 12th grade	$0.055^{***}$	$0.040^{***}$	$0.071^{***}$
	(0.015)	(0.014)	(0.018)
	Mean= $0.408$	Mean= $0.466$	Mean=0.348
Number of localities Number of observations	$89 \\ 316,322$	$89 \\ 163,262$	$89 \\ 153,060$

• Estimated effect is higher for women, especially in the higher grades.

Women were 7 percentage points (20%) more likely to graduate from high school due to school openings. 4 percentage points increase (9%) for men.

- Possible heterogeneity in the effect across different types of localities
  - Population Size (large vs small localities)
  - Distance to high school prior to high-school openings
  - Low vs high educational attainment in 1972
  - Low vs high employment rate in 1972 (proxy for socioeconomic status)
  - Low vs high traditionalism (proxy- share of young married individuals)
  - Different religions

Dependent Variable

**HS** Access

HS Access  $\times$  High pop. size

HS Access

HS Access  $\times$  High distance to school prior to local school opening

HS Access

HS Access  $\times$  Druze locality

CompletedCompletedCompleted10th Grade11th Grade12th Grade

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### Heterogeneity: short-term outcomes

Dependent Variable	Completed 10th Grade	Completed 11th Grade	Completed 12th Grade
HS Access	0.071***	0.066***	0.063***
	(0.013)	(0.015)	(0.016)
HS Access $ imes$ High pop. size	0.034	0.002	0.011
	(0.023)	(0.027)	(0.024)
HS Access	0.069***	0.049***	$0.048^{***}$
	(0.015)	(0.012)	(0.012)
${f HS} \ {f Access} \  imes \ {f High} \ {f distance}$	$0.093^{***}$	0.063**	$0.061^{**}$
to closest locality with school	(0.030)	(0.031)	(0.030)
HS Access	0.080***	$0.046^{**}$	$0.045^{**}$
	(0.025)	(0.019)	(0.018)
HS Access $\times$ Druze locality	-0.042	-0.009	-0.011
	(0.028)	(0.025)	(0.027)
• No significant heterogeneity in the effect on HS en	rollment and completion by p	oop. size, religion a	nd prior avg.
education in the locality.			

#### Dependent Variable

HS Access

HS Access  $\times$ yrs of educ  $\geq$  median (1972)

HS Access

HS Access  $\times$ employment rate $\geq$ median (1972)

HS Access

HS Access  $\times$ Share married (18-30)  $\geq$ median (1972) Mean Outcome Number of Observations Number of localities CompletedCompletedCompleted10th Grade11th Grade12th Grade

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### Heterogeneity: short-term outcomes

Dependent Variable HS Access	Completed 10th Grade 0.081*** (0.017)	Completed 11th Grade 0.057*** (0.017)	Completed 12th Grade $0.053^{***}$ (0.016)
HS Access $\times$	-0.035	-0.019	-0.018
${f yrs}  {f of}  {f educ} \geq {f median}  (1972)$	(0.029)	(0.031)	(0.033)
HS Access	0.084***	0.062***	0.055***
	(0.022)	(0.019)	(0.017)
HS Access $\times$	-0.031	-0.016	-0.006
${ m employment} \ { m rate}{\geq}{ m median} \ (1972)$	(0.038)	(0.035)	(0.033)
HS Access	$0.064^{***}$	0.066***	0.055***
	(0.015)	(0.015)	(0.015)
HS Access $\times$	0.034	-0.010	0.001
Share married $(18-30) \ge median (1972)$	(0.030)	(0.026)	(0.025)
Mean Outcome	0.563	0.439	0.400
Number of Observations	300,290	300,290	300,290
Number of localities	76	76	76
• No significant heterogeneity in effect on HS enrollment and	completion by educ	ation rate, employm	ent rate and
traditionalism			J

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# Long-term Outcomes

<b>Dependent Variable</b> Years of education $\geq 13$ (Post-secondary proxy)	All	Men	Women
Years of education $\geq 15$ (Academic degree proxy)			

Number of localities Number of observations

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Dependent Variable	All	$\mathbf{Men}$	Women	
Years of education $\geq 13$	$0.011^{**}$	0.007	$0.016^{**}$	
(Post-secondary proxy)	(0.006)	(0.007)	(0.008)	
	Mean=0.173	Mean=0.212	Mean=0.132	
Years of education $\geq 15$	0.012***	0.011**	0.013**	
(Academic degree proxy)	(0.004)	(0.005)	(0.007)	
	Mean=0.123	Mean=0.147	Mean=0.098	
Number of localities	89	89	89	
Number of observations	$316,\!322$	163,262	153,060	

• 1.1 percentage point (6%)  $\uparrow$  in post-secondary studies.

• 1.2 percentage point (10%)  $\uparrow$  in completion of 15 years of education (academic degree proxy).

# Labor market Outcomes

### Baseline Results: Labor

**Dependent Variable** Total **years** employed in ages 33-35 All

 $\mathbf{Men}$ 

Women

Total months employed in ages 33-35

Average annual wages in ages 33-35

(Log) Annual wages in ages 33-35

Number of localities Number of observations

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# Baseline Results: Labor

Dependent Variable	All	$\mathbf{Men}$	Women
Total <b>years</b> employed in ages 33-35	0.012	-0.010	$0.039^{*}$
	(0.015)	(0.023)	(0.024)
	Mean=1.543	Mean=2.120	Mean=0.936
Total months employed in ages 33-35	0.188	-0.177	0.633**
	(0.157)	(0.263)	(0.257)
	Mean=15.156	Mean=21.322	Mean=8.675
Average annual wages in ages 33-35	581	-382	$1,617^{**}$
	(687)	(1,112)	(668)
	Mean=35,213	Mean=54,983	Mean=14,433
(Log) Annual wages in ages 33-35	0.038**	0.004	0.109**
	(0.019)	(0.018)	(0.046)
	Mean=10.383	Mean=10.724	Mean=9.665
Number of localities	89	89	89
Number of observations	$316,\!322$	$163,\!262$	$153,\!060$

• For women: 0.6 increase in months of work (7%)  $\uparrow$ . 11% increase in earnings  $\uparrow$ .

For men: no significant effects. (95% confidence: Effect on employment is less that 1.6%, Effect on earnings is less than 3.2%).

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# Marriage and Fertility Outcomes

Dependent Variable	All	$\mathbf{Men}$	Women
Age of marriage			

Age of having the first child

Number of localities

Dependent Variable	All	$\mathbf{Men}$	Women	
Age of marriage	$0.146^{*}$	0.070	0.217	
	(0.086)	(0.079)	(0.136)	
	Mean = 22.922	Mean = 24.618	Mean = 21.047	
	N=279,166	N = 145,161	N = 134,005	
Age of having the first child	0.129	0.070	0.176	
	(0.088)	(0.074)	(0.138)	
	Mean=24.668	Mean = 26.374	Mean=22.844	
	N=271,461	N = 137,736	N = 133,725	
Number of localities	89	89	89	
• Increase in age at marriage by 0.1	5 years (6.5% of total ir	crease over the st	udy period)	
Increase in the age of the first child	d is similar in size but r	ot statistically sig	nificant	
- mercase in the age of the mst chil	a is similar in size but i	ior statistically sig	inicant	

**Dependent Variable** Number of children by age 20 All

Women

Men

Number of children by age 35

Number of children by age 40 Excl. 1979-1983 cohorts)

Number of localities

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# Baseline Results: Marriage and Fertility

<b>Dependent Variable</b> Number of children by age 20	$\begin{array}{c} \textbf{All} \\ \textbf{-0.023}^{*} \\ (0.013) \\ \text{Mean} = 0.221 \\ \text{N} = 316,322 \end{array}$	Men 0.001 (0.005) Mean=0.041 N=163,262	Women -0.049** (0.022) Mean=0.410 N=153,060
Number of children by age 35	$-0.120^{**}$ (0.052) Mean=3.439 N=316,322	$-0.112^{**}$ (0.046) Mean=3.209 N=163,262	$-0.128^{**}$ (0.064) Mean=3.682 N=153,060
Number of children by age 40 (Excl. 1979-1983 cohorts)	$-0.119^{**}$ (0.059) Mean=4.086 N=242,311	$-0.124^{**}$ (0.050) Mean=3.992 N=125,287	-0.116 (0.074) Mean=4.188 N=117,024
Number of localities	89	89	89

- $0.05 \downarrow$  decrease in the number of children by age 20.
- 0.12 ↓ in the number of children by age 35 and40. About 5.2% of the total decrease in the cohorts of the study (5.2 to 2.9).

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- I have also included in my study a special section for the Negev Bedouin , a special segment (10%) of the Arab population that is the most disadvantaged and are also less educated than the rest of the Arabs
- I examine the effect of recent high school openings in Bedouin communities in the Negev during the period 2007-2014
- Remarkably, I find that the effect of high school access on HS graduation for the Bedouin in 2000s-2010s is very similar to the estimated effects of the historical HS openings in 1972-1995 in the main Arab population (5-7 pp, 10-13%)

Dependent Variable	All	$\mathbf{Men}$	Women		
Any matriculation diploma	$0.067^{**}$	$0.080^{***}$	0.055		
	(0.031)	(0.025)	(0.040)		
	Mean=0.200	Mean=0.107	Mean=0.297		
Uni-eligible matriculation diploma	0.012	0.024**	0.001		
	(0.011)	(0.011)	(0.015)		
	Mean=0.121	Mean=0.054	Mean=0.191		
Number of localities	35	35	35		
Number of observations	$54,\!800$	$27,\!897$	26,903		
• Positive effect on Matriculation eligibility, but not on academic post-secondary education					

# Baseline Results: Crime (Men, Ages 16-18)

<b>Dependent Variable</b> Criminal juvenile record (all types)	No. of Criminal Records -0.248* (0.124) Mean=0.635
Property Offense	$-0.219^{**}$ (0.102) Mean=0.475
Violent Offense	-0.006 (0.012) Mean= $0.054$
Security/order offense	$-0.075^{**}$ (0.030) Mean=0.167
Number of localities Number of observations	35 27,897

• Significant decline in the number of criminal records. Particularly high decrease in property offenses.

- Differential trends or cohort effects by locality size, initial employment rate, initial traditionality.
- Controlling for locality-specific pre-trends.
- Advanced methods for staggered difference-in-differences design (Callaway and Sant'Anna, 2021; Sun and Abraham, 2021).
- Ruling out spillover effects to nearby localities
  - All produce very similar results to the baseline estimates.
- Falsification tests: placebo outcomes (5th-8th grade completion, parental education)
  - Mostly insignificant results.

- Opening a high-school in a locality boosts HS enrollment rates and also raises HS completion rates by about 5-7 percentage points (10-13%). The effect is significantly stronger in remote localities with higher distance to places with schools.
- Long term effects:
  - Positive effects on post secondary education and women's employment and earnings.
  - Negative effects on age of marriage and fertility.
- In the Bedouin localities: Significant decrease in juvenile criminal activity

- Proximity to high school matters! Geographic barriers have a significant role, especially for disadvantaged communities with low enrollment rates.
- From a global perspective, improving the supply of secondary schools in poor countries and disadvantaged communities is a positive step toward increasing global high school enrollment.
- From the Israeli perspective, improving high-school proximity for the Negev Bedouin can have potential benefits for society.
  - 30%-40% of Negev Bedouins do not have access to school in their community.

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# Thank you!

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